Formal Approaches to Decision-Making under Uncertainty

Lecture 3-1: Algorithms for DTMCs

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Reachability Probabilities

The probability to reach a goal state in $G \subseteq S$ P($\diamond G$) or P($\diamond^{\leq b} G$)



is now easily defined as

 $P(\diamond G) = P_M(\{s_0 \ s_1 \ \dots \in \text{Paths}(M) \mid \exists i: s_i \in G\})$ $P(\diamondsuit \{\text{SUNRC}\}) = 0.5 \cdot 0.5 \leftarrow 0.5 \cdot 0.4 + 0.5 \cdot 0.3 \cdot 0.5 \cdot 0.5 + \dots P(\image \{\text{Bus}\}) = 0.5$

and $P(\diamond^{\leq b} G) = P_M(\{s_0 \ s_1 \dots \in \text{Paths}(M) \mid \exists i \leq b : s_i \in G\})$

DTMC $\langle S, T, s_I \rangle$

- want to compute $P(\diamond^{\leq b} G)$ and $P(\diamond G)$

Preprocessing:

1. Make all states in G absorbing.

2. Compute state set S_0 where P($\diamond G$) from s = 0 \rightarrow graph reachability of G

Example: $G = \{3\}$ $S_0 = \{4\}$





DTMC $\langle S, T, s_I \rangle$ with probability matrix **P** – want to compute $P(\diamond^{\leq b} G)$

Let $x_i[s] \stackrel{\text{\tiny def}}{=} P(\diamond^{\leq i} G)$ from s. Then:

$$\begin{aligned} x_0[s] &= 1 \text{ if } s \in G \text{ else } 0 \\ x_i[s] &= \sum_{s'} T(s, s') \cdot x_{i-1}[s'] \end{aligned}$$

→ can formulate as matrix-vector multiplications:

$$x_i = P^i x_0$$



$$\boldsymbol{P} = \begin{pmatrix} 0 & 0.75 & 0 & 0.25 \\ 0.5 & 0 & 0.5 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{pmatrix}$$

DTMC $\langle S, T, s_I \rangle$ with probability matrix **P** – want to compute P($\diamond G$)

Let $x[s] \stackrel{\text{\tiny def}}{=} P(\diamond G)$ from s and $S_? \stackrel{\text{\tiny def}}{=} S \setminus (S_0 \cup G)$.



 $\boldsymbol{P} = \begin{pmatrix} 0 & 0.75 & 0 & 0.25 \\ 0.5 & 0 & 0.5 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{pmatrix}$

Then:

 $x[s] = 1 \text{ if } s \in G$

$$x[s] = 0 \text{ if } s \in S_0$$

 $x[s] = \sum_{s'} T(s, s') \cdot x[s] \text{ if } s \in S_?$

 \rightarrow solve linear equation system with |S| variables

DTMC $\langle S, T, s_I \rangle$ with probability matrix **P** – want to compute P($\diamond G$)

Let $x[s] \stackrel{\text{\tiny def}}{=} P(\diamond G)$ from s and $S_? \stackrel{\text{\tiny def}}{=} S \setminus (S_0 \cup G)$.



Then, if $s_I \in S_?$: $x[s] = \sum_{s' \in S_?} T(s, s') \cdot x[s]$ $+ \sum_{s_g \in G} T(s, s_g)$ $P = \begin{pmatrix} 0 & 0.75 & 0 & 0.25 \\ 0.5 & 0 & 0.5 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{pmatrix}$

 \rightarrow solve linear equation system x = Ax + b

where $A = (P(s, s'))_{s, s' \in S_{?}}$ and **b** are the one-step pr. to G

DTMC $\langle S, T, s_I \rangle$ with probability matrix **P** – want to compute P($\diamond G$)

- Solve linear equation system x = Ax + b
- 1. Direct methods:
 - Gaussian elimination, L/U decomposition, ...
- 2. Iterative methods:

Power, Jacobi, Gauss-Seidel, ...

→ iterative methods preferred due to scalability – but: inexact results



$$\boldsymbol{P} = \begin{pmatrix} 0 & 0.75 & 0 & 0.25 \\ 0.5 & 0 & 0.5 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{pmatrix}$$

Exercise I

Compute P(o { 3 }) for this DTMC by setting up and solving the linear equation system.



$$\boldsymbol{P} = \begin{pmatrix} 0 & 0.75 & 0 & 0.25 \\ 0.5 & 0 & 0.5 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{pmatrix}$$

State Elimination

Graphical variant of solving the linear equation system: eliminate states one-by-one \rightarrow the state elimination method



Exercise IIa

Use state elimination to check the correctness of the Knuth-Yao die:



(BONUS)

Extend the state elimination scheme with rewards: maintain the expected acc. reward of passing through state t





(hint: it's all about geometric series)

Value Iteration

Our P($\diamond^{\leq b} G$) algorithm – $x_0[s] = 1$ if $s \in G$ else 0 $x_i[s] = \sum_{s'} T(s, s') \cdot x_{i-1}[s]$ – is also known as value iteration:



Value Iteration

We have

$$\lim_{i \to \infty} x_i[s] = x[s]$$

 \rightarrow iterative method to approximate P($\diamond G$)

Q: When to stop?

A: Stop when
$$\forall s: |x_i[s] - x_{i-1}[s]| \le \epsilon$$

or when $\forall s: \frac{|x_i[s] - x_{i-1}[s]|}{x_i[s]} \le \epsilon$

$$x_i[s] = \sum_{s'} T(s, s') \cdot x_{i-1}[s]$$



(absolute error) (relative error)

Problem: absolute/relative error of $\leq \epsilon$ at step *i* does not guarantee $|x[s] - x_i[s]| \leq \epsilon$ (or relative variant) \rightarrow value iteration for P($\diamond G$) is **unsound**! Interval iteration

Value iteration: approximate from below

Interval iteration: approximate from below and above

Xi

true



Value Iteration

Adaptation to $R(\diamond G)$ is easy:

- 1. Precompute S_1 states where $P(\diamond G) = 1$
- 2. If $s_I \notin S_1$: return ∞ ; otherwise, iterate:



 $\begin{aligned} x_0[s] &= 0 \text{ if } s \in S_1 \text{ else } \infty \\ x_i[s] &= 0 \text{ if } s \in G \text{ else } \sum_{s'} T(s,s') \cdot (R(s,s') + x_{i-1}[s]) \end{aligned}$

Exercise III

Run value iteration for R(\diamond {3,4}) on this DTMC with an absolute-error $\epsilon = 1.2$ stopping criterion



<i>i</i> =	0	1		
$x_{i}[1]$				
$x_i[2]$				
$x_i[3]$				
$x_{i}[4]$				

Probabilistic Model Checking



Probabilistic model checking (PMC)

requires an in-memory representation of the entire DTMC – states and probabilistic transition relation

to solve the linear equation system, perform value iteration, ...

Statistical Model Checking



Statistical model checking (SMC) Monte Carlo simulation Two steps: 1. Random path generation: outcome in $\{0,1\}$ of every path = binomial random variable X_i sample mean $\overline{X} = \sum_{i=1}^{k} X_i / k$ is unbiased estimator for $P(\diamond G)$ 2. Statistical evaluation of outcomes

to ensure correct results up to error $\pm \epsilon$ with confidence $1 - \delta$

SMC: Statistical Evaluation

For $P(\diamond G)$: can use the Okamoto bound, which states that



$$\text{if } k \geq \frac{\ln\left(\frac{2}{\delta}\right)}{2 \cdot \epsilon^2}$$

Given two out of k, ϵ , δ , can compute the missing value a priori e.g. 95% confidence ($\delta = 0.05$) with $\epsilon = 0.01$ needs k = 18445 runs SMC: Sequential Testing

For $P(\diamond G) \sim x$: use e.g. Wald's sequential probability ratio test SPRT indifference **Input:** DTMC $M = \langle S, T, s_I \rangle, G \subseteq S, x \in [0, 1], d \in \mathbb{N}, \epsilon, \alpha, \beta \in (0, 1)$ **Output:** requirement satisfaction: *true*, *false*, or *unknown* type 1/11 errors 1 $p_0 := \min\{x + \epsilon, 1\}, p_1 := \max\{x - \epsilon, 0\}$ 2 $a := \log((1 - \beta)/\alpha), b := \log(\beta/(1 - \alpha))$ r := 04 repeat v := simulate(M, G, d)5 (if v = unknown then return unknown6 else if v = true then $r := r + \log p_1 - \log p_0$ 7 else $r := r + \log(1 - p_1) - \log(1 - p_0)$ 8 if $r \leq b$ then return $\sim = \geq$ // likely $P(\diamond G) \ge p_0 \ge x$ 9 // likely $P(\diamond G) \leq p_1 \leq x$ else if $r \ge a$ then return $\sim = \le$ 10

SMC for Rewards

Expected rewards:

random variables X_i are no longer binomial, but follow some unknown probability distribution

Otavioto -> confidence intervals still possible

SMC vs. PMC

Statistical model checking = random path generation plus statistical evaluation

Memory usage:

O(1) (in size of DIMC)

Runtime:

Probabilistic model checking = full state space exploration plus numeric computation

Memory usage:

Runtime:

Exercise IV

1. Let us use the Okamoto bound for SMC.

a) With k = 10000 simulation runs and desired confidence level $1 - \delta = 0.9$, what is the error ϵ that we can guarantee?

- b) Let δ and ϵ be fixed. How does k change if we halve the error, i.e. we use $\epsilon' = \epsilon/2$ instead of the original ϵ ?
- 2. Let us study a single coin flip with outcomes heads and tails. Use a real coin to estimate P(\diamond {heads}) via many "simulation runs" (= coin flips). Use the SPRT to perform only as many flips as necessary to determine whether P(\diamond {heads}) > $\frac{1}{3}$ with indifference $\epsilon = 0.05$ and $\alpha = \beta = 0.1$.

Document the steps of the SPRT algorithm as you run it.

DTMC Algorithms in Modest

Probabilistic model checking: *mcsta* Statistical model checking: *modes*

